Erasmus MC

Joint multilevel modeling of a factor analytic and a covariance regression model

Emmanuel Lesaffre

Department of Biostatistics, Erasmus MC, Rotterdam L-Biostat, KU Leuven

joint work with B. Li, L. Bruyneel and Youngjo Lee

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- Motivating data set: RN4CAST project
- Review of covariance modelling
- Multilevel Covariance Regression (MCR) model
- Multilevel Higher-Order Factor (MHOF) model



The RN4CAST project

- Registered Nurse Forecasting FP7 project (Sermeus et al., 2011)
- Nurse survey across Europe (2009-2011)
- Aim: Study the impact of system-level features of nursing care on nurse wellbeing and patient safety outcomes on burnout, ...
- Swedish data removed (no nursing unit information) & restriction to female nurses
 - \Rightarrow 21,016 nurses, 2023 nursing units, 345 hospitals, 11 countries



RN4CAST project – outcomes of interest

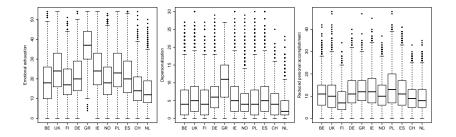
- Three dimensions of burnout
 - Emotional exhaustion (EE)
 - Depersonalization (DP)
 - Reduced personal accomplishment (PA)
- Measured using the 22-item Maslach Burnout Inventory:
 - Q: "I feel emotionally drained from my work" (EE)
 - A: 0-never; 1-a few times a year or less; ...; 6-every day
- EE (9), DP (5) & PA (8) are sum scores within each dimension



Multilevel FA and covariance model

RN4CAST project – outcomes of interest

• Distribution of burnout per country

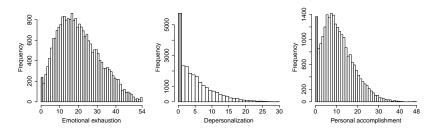




Multilevel FA and covariance model

RN4CAST project – outcomes of interest

Distribution of burnout across countries



No classical transformation to normality



RN4CAST project – covariates of interest

- Working experience (yrs): working years being a registered nurse
- Work environment: average summary of practice environment scale of nursing working index
 - Item: "Praise and recognition for a job well done"
 - Score: "Totally agree"=4, "Agree"=3, "Not agree"=2, "Totally not agree"=1
 - High values reflect a positive environment
- Teaching hosp (university hospital = 1, else = 0)
- Technical hosp ((heart/transplant) surgery present = 1, else = 0)
- Type of nursing unit (surgical = 1 or medical = 0)



RN4CAST project - covariates of interest

	Working experience(yrs)*	Work environment*	Size* [†]
Country Hospital	13.90 (9.05,18.84) 14.29 (5.05,27.76)	2.53 (2.25,2.87) 2.52 (1.71,3.26)	_ 483.60 (30,3213)
Nursing unit	13.92 (0.34,41.00)	2.54 (1.43,3.62)	11.36 (1,71)
Nurse	13.89 (0.05,50.00)	-	-

*: Mean (and range)

†: No. of beds at hospital level and No. of available nurses at nursing unit level

‡: Percentage



RN4CAST project - covariates of interest

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‡: Percentage

	Teaching	Technical	Surgery
	hospital [‡]	hospital [‡]	nursing unit [‡]
Country	_	_	
Hospital	23.77%	28.99%	
Nursing unit	-	-	49.88%
Nurse	-	-	-



RN4CAST project – research questions

• Q1: Are the means of 3 burnout dimensions associated with organizational & individual nurse characteristics?

Multivariate multilevel model for burnout means



RN4CAST project – research questions

• Q1: Are the means of 3 burnout dimensions associated with organizational & individual nurse characteristics?

Multivariate multilevel model for burnout means

• Q2: Are the variances/correlations of 3 burnout dimensions stable across hospitals, nursing units and nurses, after taking into account a rich set of confounders at different levels?

Multivariate multilevel model for burnout covariance matrix



RN4CAST project – proposed solutions

 PART I: sum scores as responses ⇒ multilevel covariance regression (MCR) model



RN4CAST project – proposed solutions

 PART I: sum scores as responses ⇒ multilevel covariance regression (MCR) model

- PART II: original 22 items as responses ⇒ multilevel higher-order factor (MHOF) model
 - = combination of MCR model with multilevel factor analytic model



PART I: The Multilevel Covariance Regression (MCR) Model

Sum scores as responses (Li et al., 2013)



- Univariate multilevel (2-level) case
- Multivariate single level case
- Multivariate multilevel (2-level) case



Univariate multilevel case:

Modeling variance with covariate x*:

$$egin{aligned} m{y}_{ij} &= m{x}_{ij}^Tm{eta} + m{u}_j + arepsilon_{ij} \ m{u}_j &\sim m{N}(0, \sigma_u^2), \quad arepsilon_{ij} \sim m{N}(0, \sigma_{arepsilon ij}^2), \quad arepsilon_{ij} ot m{u}_j \ \sigma_{arepsilon ij}^2 &=
ho(m{x}_{ij}^{*T}m{eta}^*) \end{aligned}$$

x* could be continuous or categorical



Univariate multilevel case:

• Modeling variance further with random effects:

$$\begin{aligned} y_{ij} &= \boldsymbol{x}_{ij}^{T} \boldsymbol{\beta} + u_{j} + \varepsilon_{ij} \\ u_{j} &\sim \mathcal{N}(0, \sigma_{u}^{2}), \quad \varepsilon_{ij} \sim \mathcal{N}(0, \sigma_{\varepsilon ij}^{2}), \quad \varepsilon_{ij} \bot u_{j} \\ \sigma_{\varepsilon ij}^{2} &= \rho(\boldsymbol{x}_{ij}^{*T} \boldsymbol{\beta}^{*} + u_{j}^{*}), \quad u_{j}^{*} \bot u_{j} \end{aligned}$$

- Foulley et al. (1992)
- DHGLM (Double hierarchical generalized linear model) (Lee and Nelder, 2006)



Multivariate single level case:

Multiple (p) correlated responses

 $egin{aligned} \mathbf{y}_i &= oldsymbol{B} oldsymbol{x}_i + arepsilon_i \ arepsilon_i &\sim N(oldsymbol{0}, \Sigma_arepsilon) \end{aligned}$

- Σ_{ε} : $p \times p$ residual covariance matrix
- Let Σ_{ε} depend on covariates: $\Sigma_{\varepsilon}(\boldsymbol{x}^*)$
- Problem: Positive definiteness (pd) of Σ_ε(x*)



Multivariate single level case:

- Naive solution:
 - Model each covariance element directly
 - o pd not guaranteed
- Overview of alternative solutions:
 - Logarithm transformation
 - Separation strategy
 - Modified Cholesky decomposition
 - Covariance regression



Matrix logarithm transformation of Σ_{ε} (Chiu et al., 1996):

- Definition matrix logarithmic transformation of C:
 - A symmetric, then $C = \exp(A) = \sum_{s=0}^{\infty} \frac{A^s}{s!}$ is pd \Rightarrow A := log(C)
- Property matrix logarithmic transformation of C:
 - For each pd C, \exists a A symmetric, such that $C = \exp(A)$

 \Rightarrow Let upper triangular (unconstrained) elements of *A* depend on covariates \Rightarrow **pd problem is solved**

- Interpretation: submatrix of Φ_ε ≠ log(submatrix of Σ_ε)
- Too many parameters to estimate



Separation strategy (Barnard et al., 2000):

- Separate covariance matrix into SD (diag(S)) and correlation (R) parts: $\Sigma_{\varepsilon} = diag(S) R diag(S)$
- Model each element in S with x*, but assume R constant ⇒ pd problem is solved
- Not satisfying our needs here
- Too many parameters to estimate



Modified Cholesky decomposition (Pourahmadi, 1999):

- Modified Cholesky decomposition: $T\Sigma_{\varepsilon}T^{T} = D$
- Interpretation:
 - T: conditional linear regression coefficients
 - D: conditional error variances
 - *T* and *D* can be expressed in covariates in an unconstrained manner ⇒ pd problem is solved
- But only, when there is a natural ranking of responses
- And ... too many parameters to estimate



Covariance regression (Hoff and Niu, 2012):

$$\Sigma_{\boldsymbol{x}^*} = \boldsymbol{A} + \boldsymbol{B} \boldsymbol{x}^* \boldsymbol{x}^{*T} \boldsymbol{B}^T$$

- Baseline" matrix A plus a matrix depending on *x*^{*}, *B* is the coefficient matrix of *x*^{*} ⇒ pd problem is solved
- Interpretation is intuitive: quadratic relationship in covariates
- Parsimonious representation effect of covariates



Covariance regression:

• A random-effects representation:

$$oldsymbol{y}_i = oldsymbol{B} oldsymbol{x}_i + F_i imes oldsymbol{B}^* oldsymbol{x}_i^* + arepsilon_i$$

 $arepsilon_i \sim N(oldsymbol{0}, \Sigma_{arepsilon}), \quad F_i \sim N(0, 1), \quad F_i oldsymbol{\perp} arepsilon_i$

- \Rightarrow Factor model with loadings depending on covariates
- \Rightarrow Useful for modeling



Multivariate multilevel case:

- Very few publications on this subject
- We propose a solution through a factor model
- Extension of Hoff and Niu's covariance regression model = Multilevel covariance regression (MCR) model
- An example of a hierarchical (multivariate) generalized linear model with a factor structure: **HGLM factor model**



Outline:

- Model specification
- Implied marginal distribution
- Computational approaches
- Application to RN4CAST data set



A 2-level MCR model (*i* = subject, *j* = cluster):

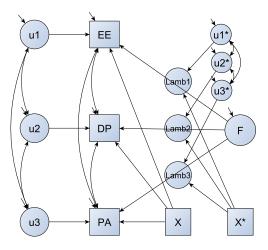
 $egin{aligned} m{y}_{ij} &= m{B}m{x}_{ij} + m{u}_j + m{\delta}_{ij} \ m{\delta}_{ij} &= m{\lambda}_{ij} F_{ij} + m{arepsilon}_{ij} \ m{\lambda}_{ij} &= m{B}^*m{x}_{ij}^* + m{u}_i^* \end{aligned}$

$$egin{aligned} oldsymbol{u}_j &\sim N(oldsymbol{0}, \Sigma_u), \quad oldsymbol{u}_j^* &\sim N(oldsymbol{0}, \Sigma_u^*) \ F_{ij} &\sim N(oldsymbol{0}, 1), \quad arepsilon_{ij} &\sim N(oldsymbol{0}, \Sigma_arepsilon) \ \delta_{ij} oldsymbol{\perp} oldsymbol{u}_j &\& \quad F_{ij} oldsymbol{\perp} arepsilon_{ij}, oldsymbol{u}_j^* \end{aligned}$$

The factor model guarantees a pd Σ_{ij} = covariance matrix of δ_{ij}



Applied to RN4CAST study:





Properties 2-level model:

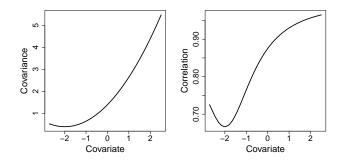
• Covariance matrix of response (conditional on random effects):

$\Sigma_{ij} = \Sigma_{\varepsilon} + (\boldsymbol{B}^* \boldsymbol{x}_{ij}^* + \boldsymbol{u}_j^*) (\boldsymbol{B}^* \boldsymbol{x}_{ij}^* + \boldsymbol{u}_j^*)^T$

- Covariates of each level can be included
- Single level case: Hoff and Niu's covariance regression model
- Easy interpretation: quadratic relationship as a function of covariates
- For 3 responses + no covariates: FA model reconstructs covariance matrix completely



Relationship between covariance/correlation and covariate:





Implied marginal distribution:

• The marginal covariance matrix of the responses is:

$$\Psi_{ij} = (\boldsymbol{B}^* \boldsymbol{x}_{ij}^*) (\boldsymbol{B}^* \boldsymbol{x}_{ij}^*)^T + \Sigma_u + \Sigma_u^* + \Sigma_{\varepsilon}$$

= $\boldsymbol{b} + \boldsymbol{a} + \boldsymbol{a}^* + \boldsymbol{c}$

- Marginal distributions of the responses are not normal
- Zero skewness for mutually independent random effects
- (excess) Kurtosis for the *q*th response:

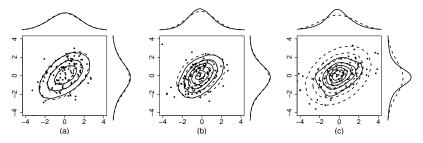
kurtosis_q =
$$\frac{6a_q^{*2} + 12a_q^*b_q}{(a_q + a_q^* + b_q + c_q)^2}$$

 a_q , a_q^* , b_q and $c_q = q$ th diagonal element of **a**, **a**^{*}, **b** and **c**



Implied marginal distribution:

• 3 bi-variate scenarios with kurtosis: (a) 0.24, (b) 1.50, (c) 3.60



Solid line (MCR model), dashed line (Gaussian model)

MCR is capable of fitting heavier-tailed distributions



- Classical likelihood approach EM algorithm (Hoff & Niu's paper, but not here)
- Bayesian approach
- h-likelihood approach



Bayesian approach:

- MCMC technique was used, since:
 - Large number of random effects & latent variables
 - Various distributions for random effects & latent variables
- Software: JAGS via R packages rjags/dclone
- Model selection: DIC and PSBF (Pseudo Bayes Factor)
- Convergence check: trace plots and Brooks-Gelman-Rubin checks
- Goodness of fit: PPC (Posterior Predictive Check) with χ^2 discrepancy function



Bayesian approach – identification issue:

No random effects in the loading part, factor part is:

 $(\beta_0^* + \beta_1^* x_{ij}^*) F_{ij}$



Bayesian approach – identification issue:

• No random effects in the loading part, factor part is:

 $(\beta_0^* + \beta_1^* x_{ij}^*) F_{ij}$

- Since $F_{ij} \sim N(0, 1)$: $(\beta_0^* + \beta_1^* x_{ij}^*) F_{ij} \iff (-\beta_0^* - \beta_1^* x_{ij}^*)(-F_{ij})$ – "flipping states"
- Different Markov chains result in symmetric (around 0) solutions
- Solution: flip negative chains over to positive



Bayesian approach - identification issue:

• With random effects in the loading part:

 $(\boldsymbol{\beta}_0^* + \boldsymbol{\beta}_1^* \boldsymbol{x}_{ij}^* + \boldsymbol{u}_j^*) \boldsymbol{F}_{ij}$



Bayesian approach - identification issue:

• With random effects in the loading part:

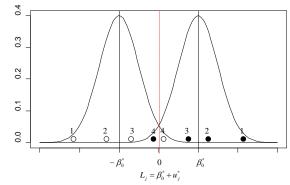
 $(\boldsymbol{\beta}_0^* + \boldsymbol{\beta}_1^* \boldsymbol{x}_{ij}^* + \boldsymbol{u}_j^*) \boldsymbol{F}_{ij}$

- Flipping states issue is more complicated:
 - $\mathbf{L}_j = \boldsymbol{\beta}_0^* + \boldsymbol{u}_j^*$
 - For cluster *j*, factor loading is $(L_j + \beta_1^* x_{ij}^*)$ or $(-L_j \beta_1^* x_{ij}^*)$
 - For uni-modal distribution of $\mathbf{L}_{j}(\boldsymbol{u}_{j}^{*})$
 - Overall mean estimate of \mathbf{L}_{i} , i.e. $\hat{\boldsymbol{\beta}}_{0}^{*}$ is close to zero
 - Σ^{*}_u is overestimated
 - Solution: Take
 - Vague prior on β^{*}₀ and β^{*}₁
 - Bi-modal distribution for L_j



Bayesian approach - identification issue:

• Solution: $\mathbf{L}_j \sim 0.5N(-\beta_0^*, \Sigma_u^*) + 0.5N(\beta_0^*, \Sigma_u^*)$



 \Rightarrow All parameters in loading part can be identified up to a sign

h-likelihood approach:

• Extended likelihood:

 $L_E(eta, \sigma, \mathbf{v} \mid \mathbf{y}, \mathbf{v}) = \prod_j \prod_i f_{eta, \sigma}(\mathbf{y}_{ij} \mid \mathbf{v}_{ij}) f_{\sigma}(\mathbf{v}_{ij})$

with $\beta = \{ B, B^* \}, \sigma = \{ \Sigma_u, \Sigma_u^* \}, v_{ij} = \{ u_j, u_j^*, F_{ij} \}$



h-likelihood approach:

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with $\beta = \{\boldsymbol{B}, \boldsymbol{B}^*\}, \sigma = \{\Sigma_u, \Sigma_u^*\}, \boldsymbol{v}_{ij} = \{\boldsymbol{u}_j, \boldsymbol{u}_j^*, \boldsymbol{F}_{ij}\}$

Lee and Nelder (1996) proposed hierarchical (h)- likelihood approach

- h-likelihood = extended likelihood when random effects combine additively with fixed effects in linear predictor
- log(L_E) is called h-likelihood
- Marginal likelihood computed by Laplace approximations
- Here h-likelihood approach generalized to HGLM factor models



Application to RN4CAST study

Outline:

- Description of data and research aims (again)
- Multilevel model for RN4CAST
- Modeling aspects
- Results: statistical and clinical



Description of data and research aims

RN4CAST data set

- 21,016 nurses, 2023 nursing units, 345 hospitals, 11 countries
- Outcomes: EE, DP, PA
- Covariates: working experience, work environment, teaching hospital, technology hospital, type of nursing unit
- Aims: Evaluate in a multi-level context relationship of covariates with
 - Means of burnout
 - Covariance matrix of burnout



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Aims: Evaluate in a multi-level context relationship of covariates with

- Means of burnout
- Covariance matrix of burnout
- Implies fitting a 3-variate 4-level model in mean and covariance
 - Here: results of Bayesian analysis
 - H-likelihood approach gave basically the same results



Multilevel model for RN4CAST

3-variate 4-level model in mean & covariance:

$$\begin{aligned} \mathbf{y}_{ijkl} &= \mathbf{B} \mathbf{x}_{ijkl} + \mathbf{u}_{jkl} + \mathbf{u}_{kl} + \mathbf{u}_{l} + \mathbf{\delta}_{ijkl} \\ \delta_{ijkl} &= \mathbf{\Lambda}_{ijkl} F_{ijkl} + \varepsilon_{ijkl}, \quad \mathbf{\Lambda}_{ijkl} &= \mathbf{B}^* \mathbf{x}_{ijkl}^* + \mathbf{u}_{jkl}^* + \mathbf{u}_{kl}^* + \mathbf{u}_{l}^* \\ \mathbf{u}_{jkl} &\sim N(\mathbf{0}, \Sigma_u), \quad \mathbf{u}_{kl} \sim N(\mathbf{0}, \Sigma_h), \quad \mathbf{u}_{l} \sim N(\mathbf{0}, \Sigma_c) \\ \mathbf{u}_{jkl}^* &\sim N(\mathbf{0}, \Sigma_u^*), \quad \mathbf{u}_{kl}^* \sim N(\mathbf{0}, \Sigma_h^*), \quad \mathbf{u}_{l}^* \sim N(\mathbf{0}, \Sigma_c^*) \\ F_{ijkl} \sim N(0, 1), \quad \varepsilon_{ijkl} \sim N(\mathbf{0}, \Sigma_c) \\ \text{All random parts independent} \end{aligned}$$



Modeling aspects

 Covariates aggregation: partition the covariate into each level (Neuhaus and Kalbfleisch, 1998)

$$\begin{aligned} x_{ijkl} &= (x_{ijkl} - \bar{x}_{jkl}) + (\bar{x}_{jkl} - \bar{x}_{kl}) + (\bar{x}_{kl} - \bar{x}_{l}) + \bar{x}_{l} \\ &= x_n + x_u + x_h + x_c \end{aligned}$$

- Non-normal burnout: apply BOS approach (Lesaffre et al., 2007)
- Missing data:
 - Missing response: treat sum scores with missing items as interval censored
 - Missing covariates: assume stochastic + jointly sample from posterior predictive distribution



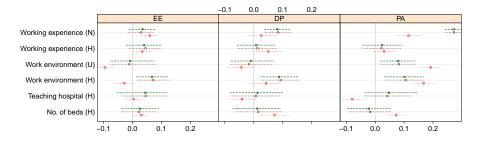


- **Statistical**: MCR model is a significant improvement (DIC, PSBF) over considering equal covariance matrices
- Clinical Mean part of burnout:
 - Longer working experience \Rightarrow less burnout at all levels
 - Better work environment \Rightarrow less burnout at nursing unit & hospital level
 - Nurses working in surgical nursing unit have more burnout
- Clinical Covariance part of burnout:
 - Correlations rather stable across models
 - Experienced nurses have a larger variance of burnout
 - Random effects: variance of burnout differs across units



Results: statistical and clinical

Fixed effect estimates in factor loadings

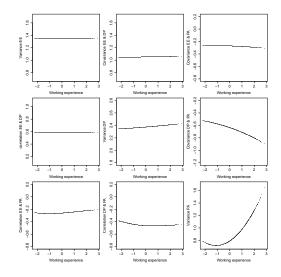


brown dashed-dotted line (our model)



Results

Impact of covariates on (co)variances and correlations





PART II: The Multilevel Higher-Order Factor (MHOF) model

Original 22 items as responses (Li et al., 2014)



- Burnout was originally measured through 22 items
- Three dimensions proposed by Maslach and Jackson (1981) were obtained from a different population
- These dimensions may be different in RN4CAST study
- Alternative: model the original 22 items directly



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- Alternative: model the original 22 items directly
- Possible analysis strategies:
 - Higher rank model:

$$\begin{split} \boldsymbol{y}_i &= \boldsymbol{B}\boldsymbol{x}_i + F_i \times \boldsymbol{B}^* \boldsymbol{x}_i^* + G_i \times \boldsymbol{B}^{**} \boldsymbol{x}_i^* + \varepsilon_i \\ \varepsilon_i &\sim N(\boldsymbol{0}, \Sigma_{\varepsilon}), \quad F_i \sim N(0, 1), \quad F_i \bot \varepsilon_i, \quad G_i \sim N(0, 1), \quad G_i \bot F_i, \quad G_i \bot \varepsilon_i \end{split}$$



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• First multilevel factor model (MFA) to find 'intrinsic' burnout dimensions (MFA), then MCR model



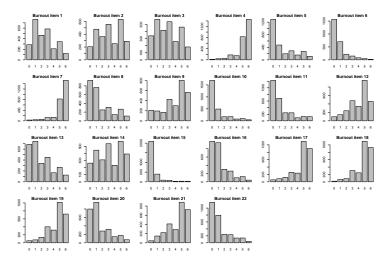
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 First multilevel factor model (MFA) to find 'intrinsic' burnout dimensions (MFA), then MCR model

• Jointly estimate MFA model and MCR model – MHOF model

Distribution of the 22 original burnout items



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Multilevel factor analytic model

- Find the latent factors underlying a group of variables in a multilevel context
- A two level MFA model is:

$$\begin{split} \mathbf{y}_{ij} &= \boldsymbol{\mu} + \mathbf{L}_B \mathbf{f}_j + \mathbf{u}_j + \mathbf{L}_W \mathbf{f}_{ij} + \varepsilon_{ij} \\ \mathbf{f}_j &\sim N(\mathbf{0}, \boldsymbol{\Sigma}_{fB}), \quad \mathbf{u}_j \sim N(\mathbf{0}, \boldsymbol{\Sigma}_u) \\ \mathbf{f}_{ij} &\sim N(\mathbf{0}, \boldsymbol{\Sigma}_{fW}), \quad \varepsilon_{ij} \sim N(\mathbf{0}, \boldsymbol{\Sigma}_{\varepsilon}) \\ i &= 1, 2, ..., n_j; \ j = 1, 2, ..., k \\ \text{All random parts independent} \end{split}$$

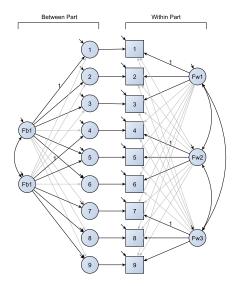
Implied covariance matrix for the MFA model is:

$$\Sigma = \boldsymbol{L}_{B} \boldsymbol{\Sigma}_{fB} \boldsymbol{L}_{B}^{T} + \boldsymbol{\Sigma}_{u} + \boldsymbol{L}_{W} \boldsymbol{\Sigma}_{fW} \boldsymbol{L}_{W}^{T} + \boldsymbol{\Sigma}_{\varepsilon}$$



Multilevel FA and covariance model

Multilevel factor analytic model





Multilevel higher-order factor model

Multilevel higher-order factor (MOHF) model =

- Is a combination of MFA model with covariance regression model
- If there are 3 intrinsic factors, then MCR model for covariance regression model is a good option
- If there are > 3 intrinsic factors, then MCR model is less optimal but can still be considered



Multilevel higher-order factor model

The MFA part:

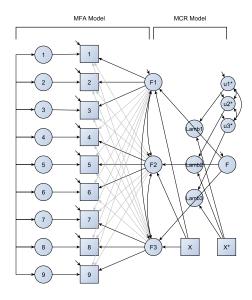
- Factor structure at the lowest level only
- Estimate the whole covariance matrix at higher levels

The MCR part:

- Use the lowest factor scores as responses
- Include covariates at each level



Model specification





Application to RN4CAST study

- Applied to Belgian part of RN4CAST study
- A 3-level MFA model based on 22 items and a 3-variate 3-level MCR model are jointly estimated
- Same modeling aspects as for MCR model
- Basically same clinical results as before



Conclusion + discussion

- MCR and MHOF inspired by clinical questions
- Alternative: SEM software (Mplus), but cannot handle this problem
- Informative priors would not change the outcome of the study
- A simulation study showed good frequentist performance of our approach



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- MCR and MHOF inspired by clinical questions
- Alternative: SEM software (Mplus), but cannot handle this problem
- Informative priors would not change the outcome of the study
- A simulation study showed good frequentist performance of our approach
- h-likelihood approach is interesting alternative + yielded same results as Bayesian approach, in a much shorter computation time:
 - Bayesian approach (rjags, dclone): 15 hours ⇔ h-likelihood approach: 3 hours
 - MCMC software is quite flexible and models can relatively easy be extended
 - h-likelihood approach: software package needed to be extended, but it is a serious competitor to INLA



That's it!

FINALLY

